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Binary fraction of O and B-type stars from LAMOST data

Feng Luo^{1,2}, Yong-Heng Zhao^{1,2}, Jiao Li^{1,2}, Yan-Jun Guo^{2,4} and Chao Liu^{2,3}

- ¹ National Astronomical Observatories, Chinese Academy of Sciences, Beijing 100101, China
- ² University of Chinese Academy of Sciences, Beijing 100049, China
- ³ Key Lab of Space Astronomy and Technology, National Astronomical Observatories, Beijing 100101, China; *liuchao@nao.cas.cn*
- ⁴ Yunnan Observatories, Chinese Academy of Sciences, Kunming 650011, China

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Abstract Binary stars play an important role in the evolution of stellar populations . The intrinsic binary fraction ($f_{\rm bin}$) of O and B-type (OB) stars in LAMOST DR5 was investigated in this work. We employed a cross-correlation approach to estimate relative radial velocities for each of the stellar spectra. The algorithm described by Sana et al. (2013) was implemented and several simulations were made to assess the performance of the approach. The binary fraction of the OB stars is estimated through comparing the unidistribution between observations and simulations with the Kolmogorov-Smirnov tests. Simulations show that it is reliable for stars most of whom have six, seven and eight repeated observations. The uncertainty of orbital parameters of binarity becomes larger when observational frequencies decrease. By adopting the fixed power exponents of $\pi = -0.45$ and $\kappa = -1$ for period and mass ratio distributions, respectively, we obtain that $f_{\rm bin} = 0.4^{+0.05}_{-0.06}$ for the samples with more than three observations. When we consider the full samples with at least two observations, the binary fraction turns out to be $0.37^{+0.03}_{-0.03}$. These two results are consistent with each other in 1σ .

Key words: binaries: spectroscopic — techniques: radial velocities — stars: early-type

1 INTRODUCTION

Binary stars play a crucial role in the evolution of stars and galaxies (Chini et al. 2012; Almeida et al. 2017). Almost a half of solar-type stars are located in binary systems (Raghavan et al. 2010; Moe & Di Stefano 2017). A nearby companion would probably affect the evolution of massive stars in binary systems (Podsiadlowski et al. 1992; Langer et al. 2008; Eldridge et al. 2011), leading to phenomena such as stellar mergers, X-ray binaries and gamma-ray bursts (Sana et al. 2012). Therefore, it is nontrivial to identify binary system from single stars and to determine the orbital parameters of the binary stars in different Galactic environments (Gao et al. 2017).

Compared the distance to earth, the separations of most binaries are too close to resolve in photometry. When orbital periods are relatively short, 1000 days for example, there are two cases that binaries can be detected through spectroscopic approaches. On one hand, spectra will appear split or contain "double lines" if the primary and secondary have near degree of luminosities (Pourbaix et al. 2004; Fernandez et al. 2017; Merle et al. 2017). Bimodal peaks would emerge in the cross-correlation function of spectra for these binary systems. On the other hand, the luminosity of secondary is much smaller than primary and it cannot contribute enough flux to the spectrum. However, it can lead to orbital motion of primary and bring distinguishable radial velocity (RV) variation (Minor 2013; Troup et al. 2016; Price-Whelan et al. 2017; Badenes et al. 2018). In this paper, we studied the latter case.

A few previous works have studied the intrinsic binary fraction $f_{\rm bin}$ in different environments in the past decades. Raghavan et al. (2010) reported that $f_{\rm bin}$ of solar-type stars is about 0.34. OB stars are considered having a larger binary fraction in some literatures. For B-type stars, Chini et al. (2012) showed that the $f_{\rm bin}$ is 0.46 ± 0.03. Dunstall et al. (2015) found that there is a relatively high $f_{\rm bin}$ which is 0.58 ± 0.11. For O-type stars, 0.42 ± 0.04 came from Aldoretta et al. (2015), 0.50±0.03 from Sota et al. (2014), 0.51±0.07 from Kobulnicky et al. (2014), 0.68±0.03 from Chini et al. (2012), and 0.69±0.09 from Sana et al. (2012) respectively. The different results of $f_{\rm bin}$ may be caused by different methods or data samples adopted in previous

Table 1 Overview of Recent Studies of Binary Fraction ofDifferent Samples

$f_{ m bin}$	Number of samples	Spectral type	Reference
0.34 ± 0.02	454	FGK	Raghavan et al. (2010)
0.46 ± 0.03	226	В	Chini et al. (2012)
0.58 ± 0.11	408	В	Dunstall et al. (2015)
0.42 ± 0.04	161	0	Aldoretta et al. (2015)
0.50 ± 0.03	194	0	Sota et al. (2014)
0.51 ± 0.07	45	0	Kobulnicky et al. (2014)
0.68 ± 0.03	243	0	Chini et al. (2012)
0.69 ± 0.09	71	0	Sana et al. (2012)

researches. These studies seem to reveal that most early type stars are residing in binary systems. The results of binary fraction studied by the above-mentioned works are listed in Table 1.

LAMOST telescope can obtain at most 4000 spectra within one observation and hence can survey the sky in a very high efficiency (Cui et al. 2012; Zhao et al. 2012). The LAMOST targets mainly on stars and its data is widely used in researches of stellar physics and the Milky Way. Its data release 5 (DR5) contains 9 027 634 spectra obtained between 2011 October 24 and 2017 June 16.

The primary goal of this study is to estimate the binary fraction of OB stars observed with LAMOST. At the same time, the distribution law of orbital parameters were explored. Observations of different epochs for a same source may obtain different radial velocities, then binaries can be discriminated from single stars from RV variation. We employ the method proposed by (Sana et al. 2013, hereafter S13), to estimate the $f_{\rm bin}$ for the LAMOST OB stars with multiple observations. LAMOST sample and data selection are described in Section 2. Details of our method and validation are presented in Section 3. The processing results of observational data are analyzed in Section 4. Discussion and conclusion are given in Sections 5 and 6, respectively.

2 DATA

2.1 Samples from LAMOST

A catalogue of OB stars from LAMOST released in 2019 (Liu et al. 2019) contains 22 901 spectra of 16 032 stars. 1715 of them have been targeted with two or more epochs by LAMOST. We discard several spectra which SNRg < 10 or have bad pixels, and leave 1680 sources for this work. Here SNRg is the signal-to-noise ratio at g-band, which can be obtained from the header recorded in the corresponding fits file of each spectrum. Each spectrum was checked by eyes, and pixels with flux equal to 0 were regarded as bad pixels.

In general, for a given binary candidate, the more observations we have, the more orbital information we can obtain. Several simulated experiments with mock data were made to assess the capability of repeated observation times to predict orbital parameters. We generated nine sets of mock samples, each set has the same sample size of 328, the same prior distributions of orbital parameters, but different observation epochs in each set of samples from 2 - 10. The simulated results show that the uncertainty of predictions of orbital parameters becomes smaller when the observational times increase.

According to our experiments, samples which have been observed only twice help little to predict the distribution of orbital parameters. Therefore, 328 sources containing 1126 spectra with at least three repeated observations were chosen for this study. The distribution of spectral type for the OB stars given by the catalog is shown in the left panel of Figure 1. Yellow bars indicate the whole samples, while red bars the subset to be used in binarity study. The distribution of time intervals for the repeated observations is shown in the middle panel of Figure 1. The right panel of Figure 1 shows the distribution of the number of observations of our samples, and it also shows that about 77.8% stars have three repeated observations.

2.2 Relative Radial Velocity

LAMOST pipeline (Luo et al. 2015) provides the Doppler shift of each spectrum, which could be used to estimate RV simply by multiplying light speed. However, the Doppler shifts in LAMOST RVs are not sufficiently precise for early type stars, since their spectra have fewer lines.

In order to get more precise RVs, we employed a cross-correlation approach to estimate the relative RVs. In general, the usage of template-matching method within a group of spectra corresponding to an observational source that observed at different time can avert the risk of inconsistency of spectral type between the spectrum to be measured and empirical template spectral library, as well as the error brought by this procedure.

The wavelength range used in the cross-correlation analysis depends on the features of OB spectra being considered within the spectral region. Wavelength range of 3900 – 5000 Å was chosen because: 1) There are absorption lines such as H_{β} , H_{γ} and H_{δ} in this range which can be used to estimate the relative RV. 2) Many OB stars are brighter at this wavelength range. The wavelength calibration is by use of arc lines, a vacuum wavelength scale is applied for it (Luo et al. 2015).

We chose the spectra with highest SNRg for each star as the template which relative radial velocity is fixed at zero. Then the rest spectra of the same star were respectively cross-correlated with the template to calculate relative RV. We defined σ_{RV} as the standard deviation of RVs for a given source, which represents the dispersion of the RVs. Results of σ_{RV} calculated by this strategy are



Fig. 1 *Left panel*: Distributions of the spectral type of the OB stars from Liu et al. (2019) and the subset for this work are presented with *yellow* and *red colors*, respectively. *Middle panel*: It displays the distribution of observational interval of the selected sub-samples. *Right panel*: It illustrates the distribution of the number of observation for the selected sub-samples.



Fig. 2 The uncertainty of RV measurement by crosscorrelation method in this paper is about 0.7 km s^{-1} . The *blue line* displays the distribution of σ_{RV} measured from cross-correlation. The *red line* shows the distribution of σ_{RV} provided by LAMOST catalog.

shown in Figure 2, blue line. As a contrast, the red line in Figure 2 indicated the σ_{RV} provided by LAMOST catalog.

2.3 Error of RV Measurement

Binaries can be detected through RV variations. However, the observed RV differences between epochs may occur only due to the measurement error. Hence, the results of radial velocity measurements for the same source are composed of both measurement error and Doppler shift caused by orbital motion within binary systems. The determination of the measurement error of RV is particularly important because it plays a significant role in the accuracy of our final analysis results.

The measurement error of RV as ε for each spectrum was determined with a Monte-Carlo simulation. For each



Fig.3 Distribution of measurement error of RV. The template spectra in each group were not included.

spectrum we draw 200 spectra with randomly added Gaussian noise. The arbitrarily drawn noise follows sigma as the following.

$$\sigma_{\text{noise}} = \sqrt{\left(\frac{f_{\text{tp}}}{snrg_{\text{la}}}\right)^2 - \left(\frac{f_{\text{tp}}}{snrg_{\text{tp}}}\right)^2},\qquad(1)$$

where $f_{\rm tp}$ is the flux of the template, $snrg_{\rm la}$ and $snrg_{\rm tp}$ are SNRg of spectra to be measured and the template, respectively. Then, each spectrum with random noise drawn by Equation (1) was cross-correlated with its corresponding template to derive the relative RVs. Finally, the uncertainty of RV of each spectrum can be derived from the standard deviation of the RVs of the 200 simulations.

The ε of each template was defined as $10^{-4} \text{ km s}^{-1}$ for the following two reasons. Firstly, to avoid the exception with a divisor of zero, and make sure that the calculations go on wheels. Secondly, it should be defined as a sufficiently small value. The ε of each spectrum was derived from the template-matching method with its corresponding template, but it will be very small for the template spectrum itself. We tested the algorithm by adopting ε as 10^{-3} and $10^{-2} \text{ km s}^{-1}$, there is no change in

the final results. The distribution of ε of all sources but the templates is shown in Figure 3.

3 METHOD

3.1 Introduction to S13 Method

 Table 2 Kinematics Orbital Parameters of Binary

Name	Comment	
P	orbital period	
q	mass ratio	
e	eccentricity	
m_1	mass of the primary	
i	angle of inclination	
ω	longitude of the periastron	
T_0	T_0 the time of periastron passage	

There are seven parameters that describe binary orbits as shown in Table 2. The RV equation (Eq. (2)) can be drawn through the geometric calculation of the motion of elliptic orbit.

$$RV = \frac{2\pi a_1 \sin i}{P\sqrt{(1-e^2)}} [\cos(\theta+\omega) + e\cos\omega] + \gamma, \quad (2)$$

where γ is the systemic velocity, or the radial velocity of the center of mass of the binary system. It relates to the motion of the whole system. The γ always equals 0 in our method since the radial velocities measured by templatematching approach within a source are relative radial velocities, which did not include the systemic velocity. a_1 is the semimajor axis of the primary star's orbit. θ is the position angle between the radius vector and a reference direction, which is a function of the time of an observational epoch. P, e, i, ω and T_0 are defined in Table 2. We thus can get a simulated RV by given a set of the seven orbital parameters and an observational time.

We implemented the algorithm described by S13. The criteria are based on the detectable obvious change in RVs. A source is deemed a spectroscopic binary star when there is at least one pair of RV satisfying

$$\frac{|v_i - v_j|}{\sqrt{\varepsilon_i^2 + \varepsilon_j^2}} > 4.0 \quad \text{and} \quad |v_i - v_j| > C \,, \qquad (3)$$

where v_i and ε_i are the RV and its error of epoch *i* for a given object. The value of *C* was conservatively adopted with 20 km s^{-1} because not only photospheric variations in supergiants can mimic variations with amplitudes of up to 20 km s^{-1} (Ritchie et al. 2009) but also wind effects of some stars (Li et al. 2016; Abdul-Masih et al. 2020). Even though, it is somewhat subjective to give an exact value to *C* as a cutoff for binary detection. However, the intrinsic binary fraction f_{bin} of real observations will be corrected through the simulation procedure described

as below. Therefore, a reasonable selection of C will not significantly affect the final result. For a source with more than three observations, several pairs of RVs may satisfy Equation (3), each pairs corresponds to its own observational time scale ΔHJD . The maximum RV variation ΔRV and the minimum time scale ΔHJD were recorded respectively in two sequences, which will be used to compare with the simulations utilizing Kolmogorov-Smirnov (KS) tests.

Accounting for observational biases due to sampling and measurement uncertainties we need to search for sets of distributions that reproduce the properties of the observations in three aspects: 1) the observed binary fraction; 2) the peak-to-peak amplitude ΔRV of the RV variations; 3) the minimum time scale ΔHJD for significant RV variation to be observed. Simulated and observed distributions are compared by means of KS tests. The binary fraction detected in the simulations will be compared to the observed fraction using a Binomial distribution (Sana et al. 2012). A global merit function (Ξ') is constructed as:

$$\Xi' = P_{\rm KS}(\Delta RV) \times P_{\rm KS}(\Delta HJD) \\ \times B\left(N_{\rm bin}, N, f_{\rm bin}^{\rm simul}\right) , \qquad (4)$$

where $f_{\rm bin}^{\rm simul}$ and $N_{\rm bin}$ are the binary fraction and the number of samples detected in simulated samples. N is the sample size. Following S13, we use power-law functions to describe the intrinsic distributions of orbital parameters, e.g., $f (\log_{10} P/d) \sim (\log_{10} P)^{\pi}$, $f(q) \sim q^{\kappa}$ and $f(e) \sim e^{\eta} (\eta \text{ is fixed at } -0.5)$. We thus can explore the distribution of Ξ' in the three-dimensional parameter space $(\pi, \kappa, f_{\rm bin})$ by using the Monte-Carlo approach.

3.2 Validation

3.2.1 Verification of the algorithm with S13 data

We used the real observational data from S13 to test our code. Following S13, we adopted the detected range of $\log_{10} P$, q and e are $\log_{10} P \in [0.15, 3.5]$, $q \in [0.1, 1.0]$ and $e \in [10^{-5}, 0.9]$, respectively. The results of our program are shown in Figure 4.

Panels in the top row are the projections of Ξ' from the multi-dimensional space onto π vs. κ , π vs. f_{bin} , and κ vs. f_{bin} . Red points indicate the peak positions. In Figure 4, the bottom panels are one-demensional Ξ' projections to π , κ , and f_{bin} , respectively.

It is worthy to note that assuming the binary mass ratio follow a power-law distribution may not be very precise (Liu 2019). However, it could tell us approximately about the relative amount of binaries with low and high mass ratio to some extent.



Fig. 4 The figure shows the results using S13 data. Panels in the top row display the projection of Ξ' onto π vs. κ , π vs. f_{bin} and κ vs. f_{bin} respectively. The *green* and *blue contour lines* indicate the 15 and 85 percentile of Ξ' . Panels in the bottom row show the projection of Ξ' onto π , κ and f_{bin} respectively. The *green* and *red dotted lines* indicate the position of 15 and 85 percentile of Ξ' .

The best-fit parameters displayed in the figure well agree with results by S13. That means we successfully implemented the algorithm proposed by S13.

3.2.2 Effect of the true $f_{\rm bin}$

More simulations were carried out to investigate how binary fraction affects the results. We generated 332 mock samples of binaries with $\pi, \kappa, \eta = (-0.45, -1, -0.5)$ (S13) and 332 samples of single stars with measurement error of each spectrum. The number of observations, observational time interval and RV measurement error of simulated binary samples are came from S13 data. The primary mass, m_1 , was randomly drawn from an initial mass function (IMF) described by Salpeter (1955) with mass range from 15 to 80 M_{\odot} . The rest parameters i, ω, T_0 are randomly drawn from the assumed distributions:

$$\cos i \sim \text{Uniform}(0, 1),$$

$$\omega \sim \text{Uniform}(0, 2\pi), \qquad (5)$$

$$T_0 \sim \text{Uniform}(0, P).$$

In order to investigate whether the adopted method can accurately estimate $f_{\rm bin}$, which is the most important parameter, we run the Monte-Carlo procedure with a parameter grid where $\pi \in [-3, 2]$, $\kappa \in [-3, 2]$. In all the following calculations we adopted the fixed $\eta = -0.5$.

The simulated samples are given with different initial $f_{\rm bin}$ ranging from 0.1 to 0.9 with a step of 0.2. The results

of $f_{\rm bin}$ estimation and predictions of π, κ are shown in Figure 5. The error bars in each panel indicate the 15 and 85 percentiles, points between top and bottom of error bars are the peak position.

Figure 5 shows that when $f_{\rm bin}$ is low, the predictions of π and κ are inaccurate. The errors of π and κ decrease and the predicted values of π and κ tend to be pinned down at around the ground true values with the increase of $f_{\rm bin}$. This is in all probability because the estimations of orbital distributions rely on the enough fraction of binary systems.

Our exercise shows that when $f_{\rm bin} > 0.3$ the estimation of π and κ is reliable. From the right panel we can see the $f_{\rm bin}$ prediction is more accurate than π and κ . This demonstrates that the method is credible at least in the $f_{\rm bin}$ estimation for the mock samples.

3.2.3 Validation with mock data mimic LAMOST samples

We then performed the third simulations to investigate whether the LAMOST samples fit the method. Mock samples mimic LAMOST OB stars with different f_{bin} from 0.1 to 0.9 with step of 0.1, were generated in the same way mentioned above but the range of m_1 was from 5 to 20 M_{\odot} . The differences are the value range of $\log_{10} P$ is [0.15, 3.0], and the detection range of π, κ are $\pi \in [-3, 3], \kappa \in [-3, 3]$. The corresponding results are shown in Figure 6.



Fig. 5 Results of simulation experiments with mock samples generated based on S13 data. The three panels display the recoveries of π , κ and f_{bin} respectively with different true values of f_{bin} , *short transverse lines* at the bottom and top of each *vertical line* indicated the position of 15 and 85 percentile, *points* between them are peak positions.



Fig. 6 Results of simulations with mock samples generated based on LAMOST OB samples. The three panels display the recoveries of π , κ and f_{bin} respectively with different true values of f_{bin} , *short transverse lines* at the bottom and top of each *vertical line* indicated the position of 15 and 85 percentile, *points* between them are peak positions.

The left panel displays the estimation of π , from which we found that there are evident asymmetry and slight overestimates in almost all samples with different f_{bin} . When $f_{\text{bin}} = 0.7, 0.8, 0.9$, the corresponding peak values of π overlapped with their corresponding 85 percentiles meaning that the probability distribution of π is highly skewed toward large values. The estimation of κ , delineated in the middle panel, shows that there are also obvious asymmetry and underestimations in all cases. Compared with S13 results we find the peak value of κ is skewed toward smaller values. But the skewness may be also caused by the limit of detection range we defined.

From the right panel we can see that $f_{\rm bin}$ estimates are roughly around the ground truth in all cases. However, the derived $f_{\rm bin}$ is slightly underestimated by around 0.1 when the real $f_{\rm bin}$ is smaller than 0.6. When the true value is larger than 0.6, the estimates are underestimated by about 0.15. The different behavior of $f_{\rm bin}$ between Figure 5 and Figure 6 is likely due to the difference of the two dataset especially their differences in observation epochs.

4 RESULT

4.1 Estimation of $f_{\rm bin}$ with Open π and κ

By means of Equation (3), 72 samples were satisfied and are regarded as binary in the 328 OB samples selected in Section 2.1. So the observed binary fraction is 72/328 = 0.22. The calculated results are shown in Figure 7, which layout is similar to Figure 4.

As shown in the top-left panel, the Ξ' range larger than 85 percentile is larger compared to S13 result. The distribution of Ξ' in π vs. κ indicates that the two parameters are correlated to some extent. It seems that Ξ' shows two peaks with irregular shape in the top-middle panel.

The $f_{\rm bin}$ and κ are anti-correlated especially when $\kappa < 0$ as shown in the top-right panel. A lower κ means that our model considers more binaries with a lower q than those with a higher q. When the model chooses smaller κ more binaries having much lower mass companion. In this case, single stars are easy to contaminate to binary samples. As



Fig. 7 Results of the LAMOST 328 OB stars with at least three observations. Panels in the top row display the projection of Ξ' onto π vs. κ , π vs. f_{bin} and κ vs. f_{bin} respectively. The *green* and *blue contour lines* indicate the 15 and 85 percentile of Ξ' . Panels in the bottom row show the projection of Ξ' onto π , κ and f_{bin} respectively. The *green* and *red dotted lines* indicate the position of 15 and 85 percentile of Ξ' .

a consequence, f_{bin} is likely larger. This can well explain the anti-correlation shown in κ vs. f_{bin} plane. The bottom panels display the marginal Ξ' of π, κ and f_{bin} . All the three distributions are much broader compared with S13 (also Fig. 4). The peak positions of π and κ are 1.4 and -1(bottom-left and bottom-middle panels in Fig. 7). A larger peak value of π means there are relative more binaries with long orbital periods than S13 samples. This may be also caused by fewer observations so that gives a partial result with little statistical significance. The estimation of κ is essentially consistent with S13 with a larger uncertainty.

It seems that the current samples, with less numbers of observations, are hard to simultaneously constrain π , κ and $f_{\rm bin}$. We then turn to only constrain $f_{\rm bin}$ with the Monte-Carlo procedure by fixing π and κ at -0.45 and -1, respectively, as concluded by S13.

4.2 Estimation of $f_{\rm bin}$ with Fixed π and κ

If π and κ are fixed, a few experiments we made show that the fluctuation of statistics is remarkable when the number of random samplings in the Monte-Carlo procedure is 100. In order to weaken the statistical fluctuation, the steps of the Monte-Carlo procedure was increased to 2×10^5 . According to Equation (2), a multi-degree-of-



Fig.8 Estimation of f_{bin} under the condition $(\pi, \kappa) = (-0.45, -1)$. The top panel: It displays the result of subset containing 328 samples. The 15 and 85 percentiles are 0.34 and 0.45 respectively, and the peak position at 0.4. The bottom panel: It shows the result of whole samples containing 1680 stars. The 15 and 85 percentiles are 0.34 and 0.4 respectively, and the peak position at 0.37.



Fig. 9 Left panel: It displays the comparison of number of observation between data from S13 and LAMOST, *blue bars* indicated LAMOST data, and *red bars* represented S13 data. Right panel: It depicts the comparison of measurement error of RV between data from S13 and LAMOST, *cyan bars* indicated LAMOST data, and *orange bars* represented S13 data.

freedom formula implies that more repeated observations were need to predict the distributions of orbital periods and mass ratio. However, the $f_{\rm bin}$ could be estimated when the parameters, π and κ , are fixed. The samples excluded in Section 2 with two observations can be used under this condition. So we calculated the total 1680 samples at the same time. The observed binary fraction directly derived by Equation (3) of total samples is 0.15. Calculated results of subset and total samples were shown in Figure 8. Here the ordinate indicated the normalized Ξ' , and $f_{\rm bin}$ is $0.4^{+0.05}_{-0.06}$ for the subset, and $0.37^{+0.03}_{-0.03}$ for the total samples. The error bar becomes narrower when the number of samples goes up, and the closer peaks show that the two results are consistent with each other. Note that, with LAMOST data the adopted S13 method tends to underestimate $f_{\rm bin}$ by about 0.15 (see Fig. 6). Therefore, the real $f_{\rm bin}$ of the OB samples may reach \sim 0.65. However, since the simulations to produce Figure 6 are under different conditions, precise systematic correction for $f_{\rm bin}$ is difficult.

5 DISCUSSION

5.1 Number of Observation

In the top panels of Figure 7, it is seen that constraints on π,κ and $f_{\rm bin}$ are not very precise compared with results of S13 shown in Figure 4. That is because seven parameters, $(P, q, e, m_1, i, \omega, T_0)$, and the time of an observational epoch, were needed to derive a RV from Equation (2). Theoretically, at least seven observations for a given source are indispensable to provide sufficient information helping to the parameters determination. However, there are 255 sources in our 328 samples with exactly three observations.

Fewer information they carried leads to larger uncertainty of the orbital parameter estimates. Number of observations of S13 and our samples are displayed in the left panel in Figure 9. Most of S13 samples have six, seven and eight repeated observations, but three, four and five in our samples. That is the main reason that our results are limited by larger uncertainty.

5.2 Measurement Error of Relative RV

While selecting sources from LAMOST, we filtered samples whose RV measurement error is larger than 20 km s^{-1} . The abandoned samples generally have spectra with low SNRg or defective pixels. The samples with larger RV measurements error were contained in S13 data and thus their RV variation may be not very accurate, see the right panel of Figure 9.

5.3 Cutoff of Orbital Period

The maximum of detection range of $\log_{10} P$, or P, obviously affect the final estimation of f_{bin} . A smaller maximum of P means the model cannot detect binaries with orbital periods longer than the maximum. Then, these binaries will be regarded as single stars, and the f_{bin} goes smaller at the same time.

The cutoff of detection range of $\log_{10} P$ of our samples is 3.0, which is smaller than 3.5 adopted by S13. So the final calculated $f_{\rm bin}$ results is smaller than S13 naturally. However, the cutoff of P is not the larger the better. Due to the influences from measurement error of RV, single stars have more risk to fall into binaries when the detection range become wider.

6 CONCLUSIONS

We applied the method introduced by S13 to estimate the properties of binary stars based on the LAMOST OB samples. We find $f_{\rm bin}$ is $0.4^{+0.05}_{-0.06}$ for the samples with at least three observations and $0.37^{+0.03}_{-0.03}$ for the total samples. Many simulations were made to investigate the capability of the S13 method. We find that because the epochs of the LAMOST observations are smaller than S13 samples, π and κ cannot be well constrained. In future, we will revisit this problem with LAMOST time-domain data so that π and κ can also be estimated.

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